# VaR and Ruin Probability

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# Introduction-VaR:

Consider an insurance risk setting:

- Let {S(t)}t≥0 denote the aggregate operating losses of a company during time (0, t].
- In VaR literature, one usually concerns the random variable S(t) for fixed time t. The value at risk at level α is defined as VaR<sub>α</sub>(S(t)) = inf{I, P(S(t) > I) < 1 − α} = inf{I, F<sub>S(t)</sub>(I) ≥ α}.
- The probability level α may assume values 0.90, 0.95, or 0.99 etc. For example, if a company's capital level is VaR<sub>0.9</sub>(S(1)), then there is 90% chance the company will be able to cover its possible operating losses next time period.



- VaR is intended to be a risk measure of financial distress over a short period of time. (Pan and Duffie, 1997)
  - In finance, the time horizon is usually a number of days. For example, the Bank for International Settlements (BIS) set p to 99% and t to ten days for purposes of measuring the adequacy of bank capital. Many firms use an overnight VaR for internal purposes.
  - ▶ In insurance, Solvency II requires a 99.5% one year VaR.
  - Notice that the time horizon in the insurance setting is much larger than that used in a bank setting, perhaps because insurance transactions are much less frequent than banking transactions.

# Criticism of VaR

- VaR ignores what happens in the tails. It specifically cuts them off. A 99% VaR calculation does not evaluate what happens in the last 1%. (Einhorn 2008)
- By ignoring the tails, VaR creates an incentive to take excessive but remote risks. (Einhorn 2008)

# Criticism of VaR-Example

- Consider underwriting two potential (annual) losses X and Y, where X takes value 1000 with p = 0.001 and zero otherwise; Y takes value 10000 with p = 0.0001 and zero otherwise. An insurer can charge a premium 2 for risk X and 10 for risk Y.
- ► The annual aggregate operating loss random variables in the two situations are S<sub>1</sub>(1) = X − 2 and S<sub>2</sub>(1) = Y − 10.
- VaR<sub>0.99</sub>(S<sub>1</sub>(1)) = −2 and VaR<sub>0.99</sub>(S<sub>2</sub>(1)) = −10. That is, you don't need any capital to support underwriting the risk.
- A firm has the incentive to take risk Y for extra profit if capital requirement is determined by VaR – remote risk is ignored by VaR.

# Criticism of VaR-Example

- A remedy for this is the use of TVaR, defined by  $TVaR_{\alpha}(S(t)) = \mathbb{E}(S(t)|S(t) > VaR_{\alpha}(S(t))).$
- For our example,

 $TVaR_{0.99}(S_1(1)) = \mathbb{E}(S_1(1)|S_1(1) > VaR_{0.99}(S(t))) = 1000,$ 

 $TVaR_{0.99}(S_2(1)) = 10,000.$ 

► This means that *Y* is riskier than *X* according to TVaR.

# Criticism of TVaR-Example

- Consider two risks, X and Y: X takes value 600 with p = 0.001 and zero otherwise; Y takes value 1000 with probability 0.0005, 200 with probability p = 0.0005 and zero otherwise.
- Suppose one may charge a premium of 2 for risk X and 5 for risk Y. Then the annual aggregate losses become  $S_1(1) = X 2$  and  $S_2(1) = Y 5$ .
- $VaR_{0.99}(S_1(1)) = -2$  and  $VaR_{0.99}(S_2(1)) = -5$ .
- $TVaR_{0.99}(S_1(1)) = TVaR_{0.99}(S_2(1)) = 600.$
- This example shows that TVaR can also ignore tail risk.



- Next, we show that infinite time horizon ruin probability is naturally a remedy for this problem.
- Instead of judging how risky it is to underwrite the risk for one year, ruin theorists ask how risky it is to continue to run the same business indefinitely.

# **Binomial Risk Model**

- ► Consider running the insurance company for t years. Assume that in each year, there is a claim with probability p or no claim with probability q = 1 - p. Assume that the annual premium is one.
- Then the aggregate operating losses at year t can be modeled by the so called compound binomial risk model (Gerber, 1988):

$$\mathcal{S}(t) = (X_1 + \cdots + X_{\mathcal{N}(t)}) - t,$$

where  $t = 1, 2, 3, \cdots$  and  $N_t$  is the number of claim in the first *t* periods.

► Ruin is the event that S(t) ≥ u for some t ≥ 1, where u is the initial surplus.



#### We consider two cases

- ► case (1): (denoted by  $S_1(t)$ ), p = 0.001 and the claim sizes  $X_i$ ,  $i = 1, 2, \cdots$  be fixed value 600.
- ► case (2): (denoted by  $S_2(t)$ ), p = 0.002 and the claim sizes  $X_i$ ,  $i = 1, 2, \cdots$  be fixed value 300.
- $VaR_{0.99}(S_1(1)) = VaR_{0.99}(S_2(1)) = -1.$
- ► Ruin probability ψ<sub>1</sub>(u) = P(sup<sub>t≥1</sub> S<sub>1</sub>(t) ≥ u), where u is the initial surplus.



- Gerber (1988) showed that  $\psi_1(0) = p\mathbb{E}(X) = 0.6$  and  $\psi_1(u) = q\psi_1(u+1) + p$ , for 1 < u < 600 and  $\psi_1(u) = q\psi_1(u+1) + p\psi_1(u+1 600)$ , for  $u \ge 600$ .
- $\psi_2(u)$  can be calculated similarly.
- Ruin probabilities as a function of initial surplus in plotted in figure 1.

# Example 1



Figure: Ruin Probability as risk measure -example 1

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- This figure shows that, when using ruin probability as the risk measure
  - $\{S_1(t)\}_{t\geq 0}$  is riskier than  $\{S_2(t)\}_{t\geq 0}$
  - If one requires that the ultimate ruin probability to be less than certain level, say 0.1, then the required initial capital can be readily determined from the graph.



#### We consider two cases

- ► case (1): (denoted by  $S_1(t)$ ), p = 0.001 and the claim sizes  $X_i$ ,  $i = 1, 2, \cdots$  be fixed value 600.
- ► case (3): (denoted by  $S_3(t)$ ), p = 0.001 and the claim sizes  $X_i$ ,  $i = 1, 2, \cdots$  take values 200 and 1000 with probability 1/2.
- ►  $VaR_{0.99}(S_1(1)) = VaR_{0.99}(S_3(1)) = -1$  and  $TVaR_{0.99}(S_1(1)) = TVaR_{0.99}(S_3(1)) = 600.$

# Example 2 continued

- ▶ In case (3), Gerber (1988) showed that  $\psi_3(0) = p\mathbb{E}(X) = 0.6$  and  $\psi_3(u) = q\psi_3(u+1) + p$ , for 1 < u < 200,  $\psi_3(u) = q\psi_3(u+1) + p\psi_3(u+1-200)$ , for 200 < u < 1200and  $\psi_3(u) = q\psi_3(u+1) + \frac{1}{2}p\psi_3(u+1-200) + \frac{1}{2}p\psi_3(u+1-1200)$ , for  $u \ge 600$ .
- Ruin probabilities as a function of initial surplus in plotted in figure (2).





Figure: Ruin Probability as risk measure-example 2

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# VaR and Ruin probability as risk measures

- VaR is a risk measure of S(t) for fixed t.
- In ruin theory literature, one usually concerns with the random variable *M*(∞) = sup{*S*(*x*), 0 ≤ *x*}.
- the infinite time horizon ruin probability is defined by ψ(u) = ℙ(M(∞) > u). That is, ruin probability is a risk measure of M(∞)
- These two measures provide different information about the risk in concern.
- Instead of judging how risky it is to bet on one trial of flipping a coin, ruin theorists ask how risky it is to continue betting on a lot of trials.

#### Brownian motion risk process

Let S(t) = −µt + σW(t) be the aggregate operating losses, where W(t) is a standard Brownian motion.

• 
$$S(t) \sim N(-\mu t, \sigma^2 t)$$
.

•  $VaR_{\rho}(S(t)) = -\mu t + \sigma t^{1/2} \Phi^{-1}(\rho).$ 

► 
$$TVaR_p(S(t)) = \mathbb{E}(S(t)|S(t) > VaR_p(S(t))) =$$
  
 $-\mu t + \sigma t^{1/2} \frac{\phi(\Phi^{-1}(p))}{1-p}.$ 

## Brownian motion risk process

- ► Infinite time horizon ruin probability concerns  $M(\infty) = \sup_{t \ge 0} \{S(t)\}.$
- $F_{M(\infty)}(y) = 1 e^{2\mu y/\sigma^2}$ , for  $\mu > 0$ .
- We next illustrate how VaR and ruin probability differ in this case.



- ► case 1 (S<sub>1</sub>(t)): μ = −1, σ = 1;
- case 2 (S<sub>2</sub>(t)): μ = -10, σ = 4.8687;
- $VaR_{0.99}S_1(1) = VaR_{0.99}S_2(1) = 1.3263.$
- Ruin probabilities as a function of initial surplus in plotted in figure (3).

## Example 3



Figure: Ruin Probability as risk measure-example 4.

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## Conclusion of the examples

- VaR and TVaR consider the short term effect of a risk.
- By looking at the long term effect of the risk, ruin probability supplement VaR and TVaR as a informative risk measure.

# VaR and Finite Time Ruin probability

- Define  $M(t) = \sup\{S(x), 0 \le x \le t\}$ . for some fixed *t*.
- The ruin probability with time horizon *t* is defined by ψ(u, t) = ℙ(M(t) > u), where u is the insurer's initial capital level.
- The surplus level to ensure that the *t* year ruin probability is less than a small probability  $1 \alpha$  is  $R_{\alpha}(S(t)) = \inf\{I, F_{M(t)}(I) \ge \alpha\} = VaR_{\alpha}(M(t)).$
- Obviously,  $R_{\alpha}(S(t)) \geq VaR_{\alpha}(S(t))$

# Analysis of the time horizon

- Is the one-year horizon used by Solvency II for insurance company too long?
- ▶ What is chance of something very bad occurs during (0, *t*)?

# Analysis of the time horizon

This question has been analyzed by Boukoudh et al. (2004), in which the authors argue that, with reasonable parameters, the interim risk (*M*(*t*)) could exceed *S*(*t*) by 40%.

## Analysis of the time horizon

• 
$$M(t) = \sup\{S(x), 0 \le x \le t\}.$$

- ► It is known that  $F_{M(t)}(y) = \Phi\left(\frac{y+\mu t}{\sigma t^{1/2}}\right) e^{-2\mu y/\sigma^2} \Phi\left(\frac{-y+\mu t}{\sigma t^{1/2}}\right)$ . See for example, page 14 of Harrison (1985).
- Notice that  $F_{S(t)}(y) = \Phi\left(\frac{y+\mu t}{\sigma t^{1/2}}\right)$ .
- With this, we may compare  $\psi(u, t) = \mathbb{P}(M(t) > u)$  with  $\mathbb{P}(S(t) > u)$ .

# Analysis of the time horizon-an approximation

- For this simple case, the joint distribution of S(t) and M(t) is known, so that the relationship between S(t) and M(t) can be analyzed. however, we next consider a rough approximation.
- Instead of investigating the relationship between S(t) and M(t), we consider S(τ) and M(τ), where τ is an exponential random variable with mean t and is independent of {S(t), t ≥ 0}.

## Analysis of the time horizon-an approximation

It is well–known that M(τ) and M(τ) − S(τ) are independent and exponentially distributed with rates

$$\omega = \frac{\mu}{\sigma^2} + \sqrt{\frac{\mu^2}{\sigma^4} + \frac{2}{\sigma^2 t}}$$

and

$$\eta = \frac{-\mu}{\sigma^2} + \sqrt{\frac{\mu^2}{\sigma^4} + \frac{2}{\sigma^2 t}}$$

respectively.

#### Analysis of the time horizon

- ► Proposition 1:  $VaR_{\alpha}(M(\tau)) \sim -\frac{\log(1-\alpha)}{\frac{\mu}{\sigma^2} + \sqrt{\frac{\mu^2}{\sigma^4} + \frac{2}{\sigma^2t}}}$
- Proposition 2: The difference E[M(t) − S(t)] = 1/η. It roughly grows with order σt<sup>1/2</sup>.



By looking at the long term effect of the risk, ruin probability supplement VaR and TVaR as a informative risk measure.



# Thank you!

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